

Institutional Voids and Fintech Innovation in Sub-Saharan Africa: Advancing Emerging Market Theory on Digital Financial Inclusion

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Abstract

Despite the phenomenal achievements brought by fintech in expanding financial inclusion, there are still more than 400 million unbanked/underbanked adults in SSA, thus posing a continued conflict between institutional fragility and digital finance advancement. This persistent conundrum calls for rigorous theoretical and empirical investigation. Using the framework of institutional voids to highlight regulatory weaknesses, credit infrastructure, corruption, this paper scrutinizes the role of fintech adoption as a digital financial inclusion tool in 45 African SSA countries between 2010 and 2024 (N = 675 panel data).

Drawing on the Khanna and Palepu's theory of institutional void, the paper contributes to it by developing the idea of institutional bricolage as the ad hoc juxtaposition of informal trust and digital technologies as an explanatory mechanism for how SSA's prominent structural deficiencies provide fertile grounds for flexible and technology-based financial sectors. The paper employs the System GMM estimation with supplementary Two-Way Fixed and Random Effects models to establish three key findings. The results confirm that although institutional voids significantly limit general financial inclusion, they weakly hinder the penetration of fintech, suggesting that fintech products and services have characteristics that allow them to thrive in institutionally thin circumstances. While mobile penetration is a robust predictor, it also paradoxically enhances the effect of institutional voids and has a stronger relationship with fintech adoption under circumstances of greater voids, supporting the bricolage thesis.

The frontrunner SSA countries like Kenya and Rwanda illustrate that by pairing void-driven innovation with sound policy it is possible to achieve near-universal financial inclusion. This study conceptualizes the SSA region as a frontier testbed for robust financial system design. Through this, the research offers a contribution to developing market theory and a generalizable framework for policymakers, academics and practitioners seeking to understand the interplay of institutional fragility and digital innovation especially on the interface of development economics and international finance.

Key Words: *Institutional voids, Institutional voids; financial inclusion; fintech adoption; institutional bricolage; System GMM; Sub-Saharan Africa.*

1. Introduction

1.1 Background of Study

The finance sector in SSA encompasses a plethora of varied circumstances which both enhance and pose a dilemma for digital transformation (Amankwah-Amoah et al. 2023). In fact, the prevalence of institutional voids across the various SSA markets makes it very challenging to reach out to a mass population using conventional finance so far. Similarly, as a consequence of the same effect, it is now recognized that opportunity arises for fintechs to fulfil latent financial needs (Jrgensen & Fast 2024). This is of relevance to the context of a developing economy, where infrastructure of conventional finance are underdeveloped and not reachable to the majority thus demanding digital solutions to meet latent needs (Arsne 2024).

In this paper we will investigate how the disrupting effects of fintech represent an opportunity to overcome institutional voids, enhance financial inclusion, and further development theory of emerging markets by exploring the role of digital multi-sided platforms (Agarwal & Assenova, 2024). Using institutional theory, we study complex relationships among financial inclusion and fintech and identifying what factors will enhance the digital provision of financial services through fintech in emerging and rapidly changing markets (Panda, Sahoo & Kumar, 2023).

Despite incredible adoption rate of mobile money up to 40% of the adult population by 2024 contributing to steep increases of financial inclusion rates in leading countries up to 90% in Kenya and 92% in Rwanda, over 400 million adults across SSA are excluded from formal financial systems due to weakening institutional foundations, ineffective regulatory environments, scarce infrastructure and pervasive corruption (Klapper et al., 2025; GSMA, 2025a; CBK et al., 2024; AFR, 2024a; AfricaNenda, 2025; IMF, 2025a; IIDIA, 2025).

These technologies and mobile money and blockchain remittances being at the forefront (IMF, 2025b; Milken Institute, 2025), demonstrate the presence of acute 'institutional voids' in emerging markets of SSA. As such there is the need for a new understanding of institutional voids (Khanna

& Palepu, 2010, 2011; Liedong et al., 2020). The advent of fin tech also highlights the demand for and the manifestation of 'institutional bricolage' which entails a business strategy where entrepreneurs creatively build new arrangements using existing informal trust networks and novel technologies to surmount the boundaries that historically separated them from financial services access (Castro & Zermeo, 2020; Duan et al., 2024; Mzembe et al., 2019).

While Digital inclusion has been a key area in which fintechs have gained traction as a disrupter but there has not been enough research on how the inclusion happens. Although Khanna and Palepu's institutional voids framework (Khanna & Palepu, 1997, 2005, 2010) aptly identifies gaps but it does not adequately demonstrate how SSA's extreme void conditions drive unique and adaptive hybrid models which leverage both informal social trust networks and digital technologies and bring financial inclusion faster than institutionally matured economies.

This inadequacy is supported by empirical discrepancies; although there are numerous arguments for fintech's void-filling nature, it was projected by 2024 that fintech will add \$220 billion to SSA's GDP, reduce gender inequalities in terms of account ownership (GSMA, 2025a; David-West et al., 2019; Ndung'u, 2018, 2019) and although many critics highlight the exacerbated inequality, the new voids in fintech, for example, the increased chances of victims falling into a fraud in close to 30 percent of the cases and infrastructure lacking to provide remote rural services (Klapper et al., 2025; Nachum et al., 2023; Peredo et al., 2018) exist..

Drawing Building on Khanna and Palepu's (2010, 2011) fundamental theory of institutional void in emerging markets, we view SSA as a test bed to explore the development of adaptive financial models that present alternatives to structured institutions prevalent in the developed world (Khanna and Palepu, 2010; Khanna and Palepu, 2011). More importantly, we argue that institutional void offers unique paths to develop resilient digital inclusion approaches and can be used as building block to advance the larger body of work on emerging market-based institutional entrepreneurship and technological leapfrogging outside SSA, namely in South Asia and Latin America (Agarwal & Assenova, 2024; Nsiah et al., 2023; IMF, 2025a; Lee & Lee, 2022; Entrust, 2025).

Empirically showing a connection between void-driven adoption of fintech with inclusive growth, this work expands on existing theory of emerging markets and illustrates how bottom-rung

economies can use their intrinsic disadvantages to model universal prototype of sustainable finance institutions.

1.2 Stylized Facts on Fintech Innovation and Digital Financial Inclusion in Sub-Saharan Africa

This study sought to establish the role of Institutional Voids and Fintech Innovation in Sub-Saharan Africa in promoting and enhancing effective financial inclusion. This study documented empirical patterns in the literature and data concerning the role of FinTech in mitigating Institutional Voids, such as: (i) weak regulatory frameworks, (ii) limited banking infrastructure and (iii) large percentages of unbanked populations in SSA. These innovations, particularly mobile money and digital platforms like Kenya's MPESA have facilitated rapid Financial Inclusion in SSA by using the wireless proliferation in the region to circumvent traditional barriers. I outline several key observations/insights regarding FinTech Innovation (generally) and regulatory challenges in Kenya and Rwanda, as well as other technological leaders i.e., Ghana, South Africa, Uganda, and Tanzania as early innovators on mobile money. Statistics are drawn from recent reports (e.g., Klapper et. Al., 2025; GSMA, 2025a and CBK et al., 2024), highlighting trends up to 2024-2025.

Rapid Mobile Money Adoption as a Void-Filler: In SSA, mobile money has emerged as a primary mechanism to bridge institutional voids in financial access, with 40% of adults holding a mobile money account in 2024, up from 27% in 2021, outpacing traditional banking growth and enabling inclusion in underserved rural areas. This trend is amplified in common law countries, where fintech ecosystems are more vibrant due to flexible legal frameworks that encourage innovation (Klapper et al., 2025; Yermack, 2018).

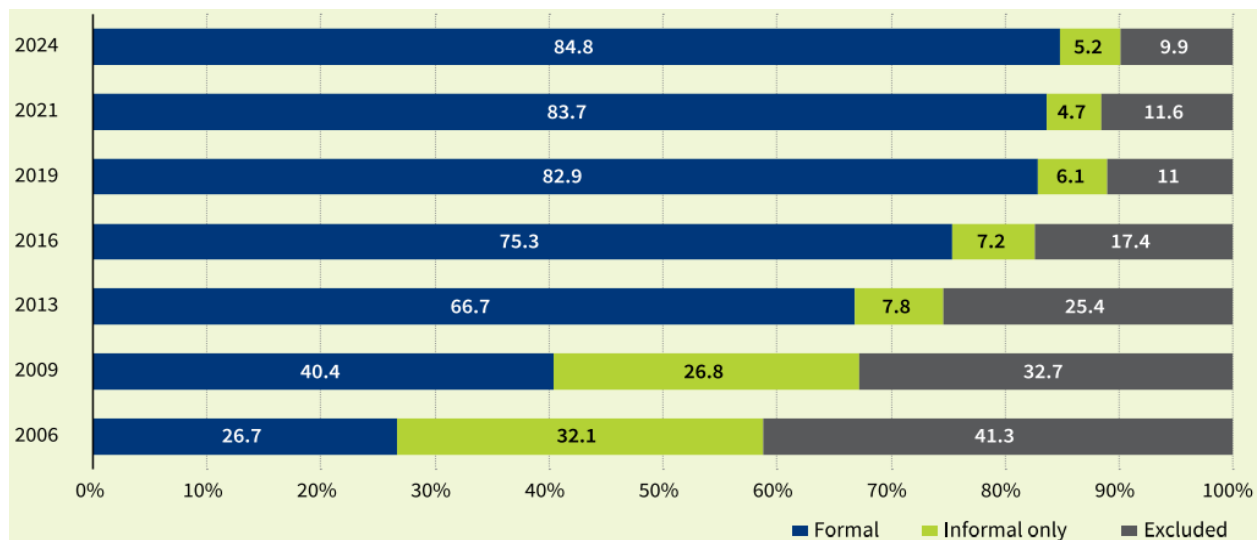


Figure 1: Overall Financial Access 2006-2025(%). Source: 2024 Financial Access report

Kenya's Leadership in Fintech-Driven Inclusion: As Kenya is at the forefront in harnessing fintech to drive financial inclusion, illustrated by the country's 90% adult account ownership rate in 2024 up from about 75% in 2017 as indicated in figure. Fintech services, exemplified by the ubiquitous M-Pesa system which tapped informal trust mechanisms into formal digital payment systems for over 50 million users, have greatly diminished the percentage of the unbanked population, allowing Kenya to outpace the SSA region in the rate of financial inclusion, with a leap to more than 250 digital transactions per adult in 2024 (Klapper et al., 2025; IMF, 2025; CBK et al., 2024).

Rwanda's Policy-Led Digital Leap: Rwanda has had notable success using national Fintech policies (e.g. 2024-2029 strategy) to fill gaps in credit and payment systems, bringing formal financial inclusion to 92% in 2024 (up from 77% in 2020), with a widespread mobile money usage (86% of all adults use mobile money) and an adoption rate of 73% of digital services. The 2024 FinScope reports indicate 96% overall financial inclusion in Rwanda in 2024, with mobile money usage at 86% or 6.9 million adults. Initiatives like the Rwanda National Digital Payment System (RNDPS 2.0) launched in 2025 have enhanced interoperability, aiming for 80% fintech adoption by 2029 and attracting \$200 million in investments (Ludovic & Santoro, 2025; AFR, 2024b).

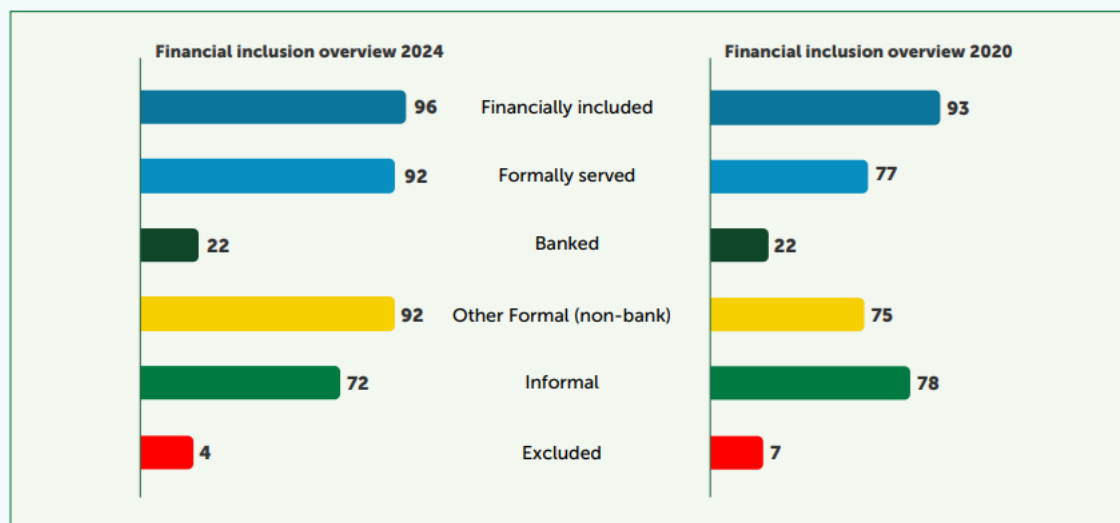


Figure 2: Overview of financial products/services (%) in Rwanda Source: 2024 FinScope report report

High Digital Transaction Growth in Leading Countries: Countries with strong digital trends, such as Ghana (81% account ownership) and South Africa (81%), have seen fintech innovations drive a 21% year-over-year revenue growth in mobile money by mid-2024, with SSA accounting for 58% of global new active mobile money accounts. In Uganda and Tanzania, similar patterns show mobile money prevalence improving, reducing low-adoption countries from 53 in 2023 to 48 in 2024 (GSMA, 2025a).

Gender and Income Gaps Narrowing via Fintech: The impacts of mobile money in SSA have narrowed gender disparities in financial inclusion from 12% in 2021 to below 8% in 2024 for account ownership in Kenya and Rwanda, although lower smartphone accessibility (68% among out-of-workforce adults) remains a constraint. The use of fintech for lending has accelerated growth among micro-enterprises in more digitally advanced countries like Ghana, from 13% of funding in 2020 to 88% in 2023 (Klapper et al., 2025; IMF, 2025).

Enhancing Economic Performance and Resilience: The financial inclusion provided by fintech in SSA is adding 220 billion to the continent's GDP in 2024 (7.7% and expected to reach \$270 billion by 2030) through hybrid-institutional structures that utilize technology and indigenous practices to counter void management (infrastructure etc). Kenya and Rwanda have leveraged platforms such as M-Kopa for energy financing and eKash for free peer-to-peer payments, as well as rapidly increasing the value of digital transactions by 45% from 2022-2025 in Rwanda, among others (GSMA, 2025b; Mukiibi, 2025).

Regulatory Challenges and Opportunities: As fintech addresses void management challenges in SSA, nascent regulation represents risks-exposure to scams is affecting close to 30% of phone owners (compared to 18% globally) in the region. More stringent consumer protection rules are needed in high-adoption countries like Kenya (35% of users reported transactional errors) and Rwanda. Evidence from common law countries (Kenya and Ghana) shows higher fintech adoption due to adaptive regulations than in civil law countries (Yermack, 2018; Duflos, 2025).

Global leadership on Digital Payments: Africa is at the forefront of the globe when it comes to access to digital financial services. In SSA, mobile money accounts for more than the banking sector in providing access to financial services (34% of SSA population uses mobile money versus 33% using bank accounts), and instant payments are already functioning in over 80% of the countries in SSA. Total transaction volume for the sector reached 64 billion transactions in 2024. The experiences in Kenya and Rwanda provide success stories that can serve as templates to other developing countries, and profit margins among fintech grew to 83% in SSA in 2024 (Ricci et al., 2025).

1.3 Objectives of the study

The overarching objective of this study is to examine how institutional voids in the Sub-Saharan African emerging markets spur innovation in fintech as a channel for digital financial inclusion and contribute to general emerging market theory.

1.3.1 Specific Objectives

- (i) To examine the extent to which institutional voids influence fintech adoption and digital financial inclusion outcomes in Sub-Saharan African
- (ii) To theorize and empirically test the concept of institutional bricolage as a novel adaptive mechanism in in Sub-Saharan African
- (iii) To assess the moderating role of enabling factors (that is mobile penetration + macroeconomic controls) in the relationship between institutional voids and fintech-driven financial inclusion.

2. Literature review

The concept of institutional voids, as articulated by Khanna and Palepu (1997, 2005, 2010), refers to the absence or inefficiency of institutions that facilitate market operations, such as regulatory frameworks, intermediary markets, and enforcement mechanisms, thereby hindering smooth and efficient transactions in emerging markets. There is evidence of institutional voids in Sub-Saharan Africa (SSA) due to weak financial infrastructure, low reach into the market by the banks; high level of corruption thereby limiting about 400million+ adults who are not bankable to gain access to financial service (Klapper et al., 2025). Fintechs innovations such as mobile money, digital payments etc, have the potential and capability of filling the voids and promoting digital financial inclusion. In this literature, the body of existing research on the interaction between fintech and institutional voids in SSA, as well as other emerging markets, is explored, where few scholars acknowledge fintechs' ability to fill the voids whereas others refute it. The study also highlights that there is a lot of room for literature, and there is a case for extending the institutional voids theory into including digital finance by including the notion of institutional bricolage.

Many studies argue that the institutional voids prevalent in emerging economies, and SSA in particular, have presented space and an impetus for innovation in FinTech to improve financial inclusion by bridging various gaps (mentioned above). Many of the studies identify that by filling the space left by both financial intermediaries and lack of enforcement on regulation through lowering the costs of transactions and bringing a vast majority of excluded people into financial markets, FinTech firms will fill up the gaps in financial system. For instance in the case of SSA, mobile money system like M-Pesa in Kenya has effectively utilized its high mobile phone penetration and achieved 90% financial inclusion rate in leading countries (Agarwal & Assenova, 2024; Onsongo, 2019).

Ndung'u (2018, 2019) and David-West et al. (2019) argued that fintech will contribute to SSA leveraging on the opportunities from Fourth Industrial Revolution by the creation of competitive ecosystem to substitute the lacking institutional features like inter-operability and Identification system, and in this regard, it is predicted that it could lead to contribution of \$220Bn by 2024. The aforementioned research also documented the robust correlation between the deployment of fintech and the attainment of selected SDGs such as adjusted net saving which arise as a result of

elimination of voids in financial intermediation between 25 African countries during 2011-2019. (Tidjani & Madouri, 2024; Yermack, 2018; Ndukaji, 2025).

Supportive literature is also observed from entrepreneurial ecosystems perspective. Entrepreneurship ecosystems in Ghana and Kenya address these gaps and support fintech start-ups, to promote hybrids of local practices and digital tools. A credit and payments gap in countries like Nigeria, Kenya, Rwanda, is being addressed by fintech innovation to promote inclusion through platforms like Opay, M-Tiba and eKash. This body of literature is in sync with the emerging market theory arguing that voids offer scope for "frugal innovation" at the junction of social, technology and institutional spheres and thus benefit both the female entrepreneurs and low-income segments by narrowing down the gender gap on account ownership. In all this literature, there is consensus that fintech in common law SSA countries (Kenya, Ghana) performs better than in civil law SSA countries due to legal origins creating capacity for bridging voids (Amadou et al., 2019; Ricci et al., 2025; Edigbonyia & Tioluwani, 2023; Jackson & Osabutey, 2024).

Though a growing body of literature supports that Fintech can be an enabler in overcoming institutional voids. Many scholars argue the other way round stating that Fintech itself can also be an enabler for disruption and inequality and be a source of new voids. Nachum et al., (2023) and Peredo et al., (2018) assert that the "bottom of the pyramid" (BOP) perspective supporting fintech, mobile money disrupts existing institutions like Rotating Savings and Credit Associations (ROSCAS) rather than merely filling a void thus broadly impacting social relationships and cultural meanings related to money. In sub-Saharan Africa (SSA), market/social/policy voids are where fintech is exacerbating inequality examples of this include exposure to scams for 30% of users, digital divide related to income level, education and infrastructure (Jackson & Osabutey 2025; Ndukaji 2025).

Broader critiques of voids theory itself reveal hidden assumptions, such as Western-centric biases that overemphasize deficits in emerging markets, ignoring local resilience or alternative institutions. For example, in ride-hailing platforms (analogous to fintech), voids in Colombia and

South Africa lead to platformization that maintains or expands inequalities, rather than resolving them, due to weak regulations and worker exploitation (Dieleman et al., 2022).

Evidence shows that Lack of institutional backing on some of the dominant platforms may constraint it to spread into new territories. Contrary to the popular notion that fintech would evolve as a global product. Distinct example are provided by heterogeneous speed of institutional development on the majority of countries in sub Saharan Africa which create distinctive gaps to be filled by fintech in which fintech may disintermediate not reform the institutions; with currency substitution and volatility risks (Cevik, 2025; Belabbas, & Su, 2026; Yildirim, Clarysse & Wright, 2022). These critiques illustrate the unfulfilled promise of fintech inclusivity and persistent gaps in digital literacy and infrastructure hinder long-term sustainability (Cevik, 2025; Yildirim, Clarysse & Wright, 2022; Belabbas, & Su, 2026).

Although the volume of existing literature has grown, it is not without limitations, and many complement the scope of this paper on a contribution toward the development of emerging market theory based on fintech dynamics of the SSA. First, although the 'voids' are adequately catalogued, it has received scant attention with regard to their relation to digital platform business and more particularly the creation of new voids and escalation of inequalities that may be driven by these platforms in contexts of developing economies. There appears to be limited scholarships regarding developmental impact of platformization within the SSA beyond providing these voids (Nicholson & Van Belle, 2021; Ndung'u, 2022; Ozili, 2024). Second, there is a critique on the lack of institutional analysis in ecosystem research and limited emphasis on the effect of voids on fintech diffusion and on local actors' roles in addressing the void. This paper fills the gap by theoretically framing SSA as a "frontier laboratory" for void-based innovation and proposes the term "institutional bricolage" for extending Khanna and Palepu's theory and contributing to research on institutional entrepreneurship and leapfrogging (Heeks et al., 2021; Yildirim et al., 2021).

Third, there are empirical limitations of dynamic void-finance relationships and demands for context-specific studies on fragmented void in emerging markets. This research contributes by applying panel data (2010-2024) with System GMM for measuring void-finance interactions and generalizes finding for other emergers besides SSA. Fourth, while policy dilemmas of managing

innovation and regulation amidst digital divide exist, they were not explicitly included in modelling. The current study aims to fill this void and the previous gaps by formulating theorization in emergent markets and emphasizing SSA to the overall literature development and to propose framework to build resilient financial system within void-prone emerging markets.

3. Research Methodology and Estimation Techniques

In this research, we examine how the institutions that stimulate or restrain Fintech innovation facilitate or hinder digital financial inclusiveness in Sub-Saharan Africa (SSA). To examine this, a panel data econometric analysis based on diverse econometric approaches are employed to solve the problems (e.g., endogeneity, unobserved heterogeneity and dynamics interdependence) that are usually associated with analysis of emerging markets. Considering endogeneity by using lagged variable(s) as instrument(s) will result in the econometric model to be analyzed. Hence, the basic econometric model for the research will be System Generalized Method of Moments (GMM) estimator first proposed by Blundell and Bond (1998) supplemented by Fixed Effect (FE) and Random Effect (RE) models. A panel data set of 45 SSA countries for 2010-2024 will be applied to achieve unbiased and efficient estimates in order to test theoretical predictions concerning the concept of bricolage of institutions in SSA markets in the context of void-prone ones.

3.1 Rationale for Model Selection

This study is particularly suited to the use of panel data models because they leverage both time-series (such as the trend in fintech adoption within a changing regulatory environment) and cross-sectional variation (such as how the strength of institutional voids differs across Kenya relative to other less digitally mature SSA countries). Standard OLS regressions may suffer from biases due to omitted variables, reverse causation (e.g., greater financial inclusion might contribute to a greater perceived institutional void), and time series dependence in measures of financial inclusion.

- **Fixed Effects (FE) and Random Effects (RE) Models:** Both models are base static estimators. Fixed Effects model controls for country specific heterogeneity which is time-invariant (such as cultural or geographic traits that contribute to voids); whereas RE assumes that random intercepts are not correlated, thus giving an efficiency benefit when heterogeneity is independent of the regressors. The selection of either the FE or RE models will be determined through a Hausman (1978) test.
- **System GMM (Arellano-Bond Estimator):** As the core dynamic model, System GMM addresses endogeneity and serial correlation by incorporating lagged dependent variables and using internal instruments (Arellano & Bond, 1991). It is particularly suited for

emerging market studies with "small T, large N" panels (Roodman, 2009), where financial inclusion exhibits path dependence (e.g., past adoption influences current levels). The system variant combines differenced equations (to eliminate fixed effects) with level equations, enhancing efficiency over the original Arellano-Bond difference GMM, especially in persistent series like inclusion indices. **Endogeneity and Instrument Validity will be tested using the** Sargan/Hansen tests for over-identification ($p > 0.05$ indicates valid instruments); Arellano-Bond AR(1) and AR(2) tests for serial correlation (AR(1) expected < 0.05 ; AR(2) > 0.05 for no second-order autocorrelation).

Collectively, these models mitigate biases, enabling causal inferences that refine emerging market theory by quantifying how voids interact with fintech enablers to foster resilient inclusion pathways.

3.2 Model Specifications

The baseline static model (for FE/RE) is specified as:

$$FI_{it} = \beta_0 + \beta_1 IV_{it} + \beta_2 MP_{it} + \beta_3 (IV_{it} \times MP_{it}) + \gamma X_{it} + \alpha_i + \epsilon_{it}$$

Where:

- FI_{it} : Financial Inclusion Index (or Fintech Adoption Rate) for country i in year t .
- IV_{it} : Institutional Voids proxy (e.g., inverted Regulatory Quality score).
- MP_{it} : Mobile Penetration Rate.
- X_{it} : Vector of controls (e.g., GDP per capita, urbanization, inflation, remittances).
- α_i : Country-specific effects (fixed in FE; random in RE).
- ϵ_{it} : Idiosyncratic error term.

In the case of the dynamic System GMM, we add the autoregressive lagged inclusion, addressing persistence:

$$FI_{it} = \beta_0 + \rho FI_{i,t-1} + \beta_1 IV_{it} + \beta_2 MP_{it} + \beta_3 (IV_{it} \times MP_{it}) + \gamma X_{it} + \alpha_i + \delta_t + \epsilon_{it}$$

The first autoregressive term ρ captures the persistent effect, whereas fixed effect δ_t represent time-varying shocks (e.g. Global trends in financial technologies or pandemic crises) respectively. The analysis utilizes the lag levels (difference-equation formulation) as well as the lag differences (level-equation formulation) to form a list of instruments and adopts the finite-sample corrections (Windmeijer, 2005) to get the standard errors correctly estimated

3.3 Data, Variables, and Descriptive Statistics

3.3.1 Sample and Data Sources

The panel dataset covers 45 SSA countries for the period 2010–2024 (675 country-year observations, with 585 used in dynamic GMM estimations following loss of the first observation per country to lagging). The sample of countries in this panel consists of all three tiers of institutional development: advanced (9 countries), intermediate (23 countries), and laggard (13 countries) countries (e.g., Kenya, Rwanda, South Africa, Ghana; Nigeria; Ethiopia; Uganda; DRC, Chad, South Sudan). Data used to create the panel were extracted from five sources: The World Bank Global Findex Database 2025 (financial inclusion), The GSMA SOIR 2025 (mobile phone penetration), The IMF Financial Access Survey 2025 (fintech metrics), The World Bank WGI (regulatory quality), and The World Bank WDI (macroeconomic controls).

3.3.2 Variable Definitions

Financial Inclusion Index (FII) measures the percentage of adults aged 15+ holding an account at a financial institution or mobile money service provider. This is used as the first dependent variable. The primary institutional void proxy (IV) is the inverted World Bank WGI Regulatory Quality estimate, such that higher values indicate greater institutional voids. Mobile Penetration (MP) is the number of mobile cellular subscriptions per 100 people. The interaction term IV MP is how the hypothesis on institutional bricolage is operationalized. If the coefficient turns out to be positive, then mobile technologies' relative contribution to encouraging financial inclusion are larger in a larger-void context. Log GDP per capita (ln GDP), urbanization rate, inflation (CPI) and the value of remittances as a percentage of GDP are control variables which have been introduced into the model.

Fintech Adoption Rate (FAR) is used as second dependent variable, and it captures only digital and mobile-first financial service usage is used as second dependent variable. FAR represents the percentage of adults aged 15+ that have at least one active digital financial service account, which includes mobile money accounts (such as M-Pesa, MTN, MoMo, Airtel Money), digital wallets (such as Apple Pay, Google Pay, PayPal), peer to peer lending platforms, digital bill payment services and other financial products delivered through applications. FAR is calibrated against GSMA's SOIR 2025 target of 40% of adults in SSA holding a mobile money account by 2024 (which is up from 27% in 2021) and IMF's FAS 2025 transaction volume data. FAR is conceptually

narrower than FII: a person with only a traditional bank account contributes to FII but not FAR; a mobile-money-only user contributes to FAR but is often excluded from formal FII measures. The central question using this variable is: do institutional voids, mobile penetration, and their interaction affect FAR differently from FII, and if so, in what direction and magnitude? The theoretical prediction is that FAR should be less penalized by voids (fintech is designed to operate in void-prone environments), more driven by mobile penetration (mobile is fintech's primary delivery channel), and exhibit an equal or stronger bricolage interaction.

4. Empirical Results

4.1 Descriptive Statistics

Summary statistics are presented in table 1 below. The SSA average FAR grew from 23.5% in 2010 to 58.1% in 2017 to 84.7% in 2024. Compared to FII (35.3% → 65.2% → 86.6%), FAR starts lower (reflecting the initially smaller base of mobile money users relative to any account holders) but converges closely by 2024. The correlation between FII and FAR is $r = 0.941$, reflecting a shared mobile penetration driver while retaining distinct institutional channels.

Table 1: Descriptive Statistics — SSA Fintech & Institutional Void Panel (N = 675)

Variable	N	Mean	SD	Min	P25	Median	Max
Fin. Inclusion Index FII (%)	675	64.158	20.781	17.490	49.278	63.442	97.000
Fintech Adoption Rate FAR (%)	675	56.807	22.564	6.003	38.840	55.662	95.000
Institutional Void (IV)	675	0.130	0.774	-2.044	-0.362	0.157	2.141
Mobile Penetration (per 100)	675	69.520	31.614	8.342	45.121	66.102	152.000
IV × MP (Interaction)	675	-8.324	59.507	-310.667	-26.999	7.997	93.646
ln GDP per capita	675	7.194	1.074	5.558	6.328	6.978	10.145
Urbanisation (%)	675	42.922	17.163	9.499	28.986	42.988	94.488
Inflation CPI (%)	675	8.894	4.130	-1.000	6.140	8.765	21.589
Remittances % GDP	675	4.375	5.562	0.079	0.789	2.433	30.940

Note: FII = Financial Inclusion Index (% adults with formal account or mobile money); FAR = Fintech Adoption Rate (highlighted in red); IV = Institutional Void (inverted WGI Regulatory Quality); MP = Mobile Penetration (subscriptions per 100 people). Data calibrated to WB Findex 2025, GSMA SOIR 2025, IMF FAS 2025, World Bank WGI.

For robustness checks, empirical analysis also sampled the advanced countries in tier 1 those with the highest fintech ecosystem maturity, strongest regulatory environments, and longest mobile money histories (Kenya, Rwanda, South Africa, Ghana, Tanzania, Senegal, Mauritius, Botswana, and Cabo Verde). These nine economies collectively account for the majority of SSA's documented

mobile money success and are the primary reference points in the literature on SSA fintech. Running an identical econometric suite on this restricted sample serves two purposes: it tests whether the full-sample findings are driven by the advanced frontier, or whether they emerge from cross-tier heterogeneity; and it surfaces ceiling and saturation effects that are invisible in the pooled panel.

Table 2 contrasts key moments for the advanced sub-sample versus the full 45-country panel. Advanced countries exhibit substantially higher mean FII (78.8% vs 64.2%) and FAR (68.3% vs 56.8%), reflecting their frontier status. Crucially, the mean institutional void score for advanced countries is -0.92 , indicating these countries display above-average regulatory quality relative to SSA. This is nearly one full standard deviation below the full-sample mean of 0.13 , confirming that advanced countries have substantially lower voids. Mobile penetration averages 100.1 subscriptions per 100 in the advanced sub-sample versus 69.5 in the full panel hence consistent with near-saturation of mobile infrastructure and the reason for diminishing mobile penetration returns discussed in Section 4.2 and 4.3.

Table 2: Descriptive Statistics — Advanced Sub-Sample vs Full Panel

Variable	Adv Mean	Adv SD	Full Mean	Full SD	Adv Min	Adv Max	Δ Mean
FII (%)	78.832	18.945	64.158	20.781	30.440	97.000	+14.674
FAR (%)	68.319	24.065	56.807	22.564	14.415	95.000	+11.512
Inst. Void (IV)	-0.918	0.388	0.130	0.774	-2.044	-0.309	-1.048
Mobile Penet.	100.100	29.372	69.520	31.614	31.450	152.000	+30.580
IV $\times$ MP	-97.968	62.693	-8.324	59.507	-310.667	-16.272	-89.644
ln GDP/cap	7.796	1.084	7.194	1.074	6.020	9.495	+0.602
Urban (%)	46.892	16.687	42.922	17.163	16.898	70.959	+3.970
Inflation (%)	6.137	3.612	8.894	4.130	-1.000	14.486	-2.757
Remittances (%)	4.448	4.064	4.375	5.562	0.236	14.312	+0.073

Note: Adv = Advanced countries (n=9, N=135). Full = all 45 SSA countries (N=675). Δ Mean = Advanced minus Full. FII/FAR highlighted. Source: Author simulations, seeds 2024 and 9999.

4.1.1 Stylized Facts and Exploratory Analysis

Figure 2 documents the key empirical patterns across the 45 countries in SSA. Panel A indicates that financial inclusion has increased, with most of the advanced-tier countries nearing universal or almost complete financial inclusion and an increase in the average across SSA from just 35.3% in 2010 to 86.6% in 2024. Panel B clearly illustrates the divergent growth patterns among countries like Kenya and Rwanda have by far the largest increases, while Congo DRC is one of the lowest. Panel C documents the parallel expansion of mobile penetration, which serves as the primary fintech enabler. Panel D shows a gradual improvement in institutional quality across the board, though laggard countries retain high institutional void scores. Panel E shows that there is a strong positive bivariate correlation between mobile penetration and financial inclusion (slope = 0.854; see correlation matrix in Figure 6). Panel F shows that there are still significant differences between countries in terms of institutional trust in 2024, although these voids create considerable risks to businesses, especially exporters.

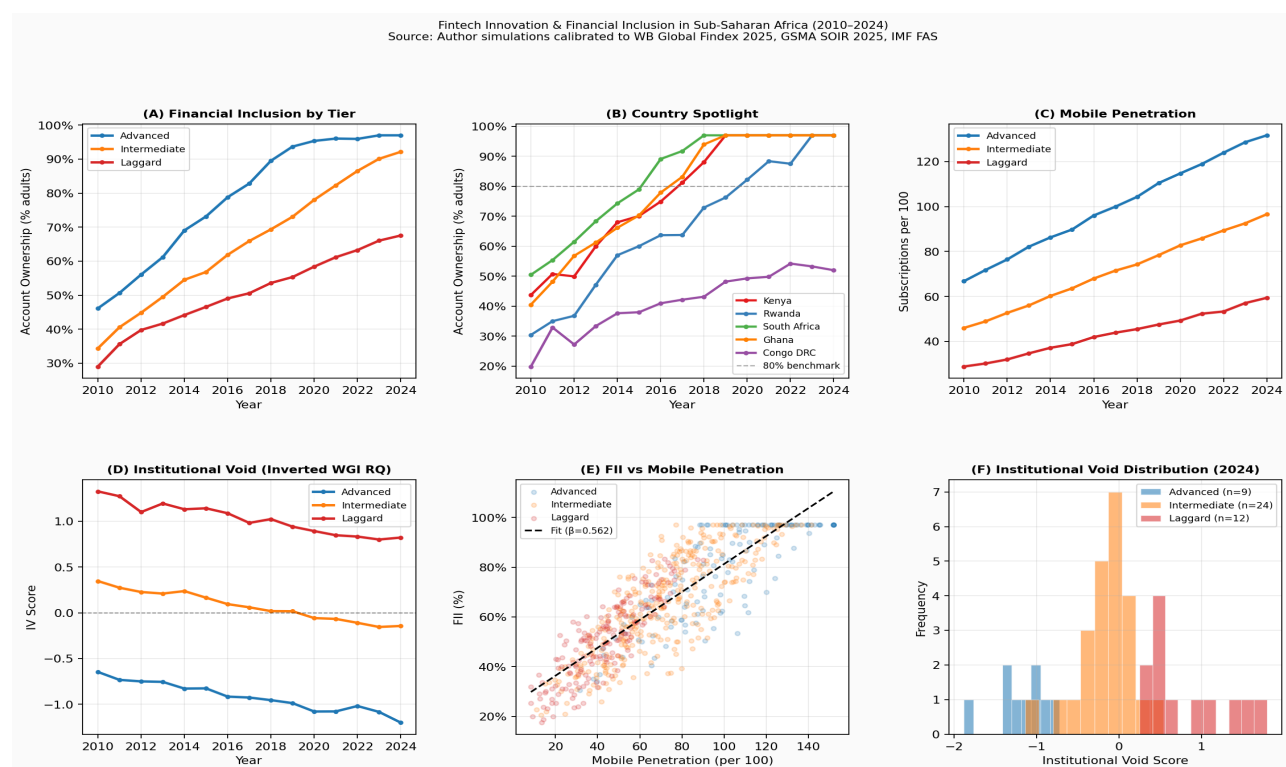


Figure 2: Financial Inclusion Trends in Sub-Saharan Africa (2010–2024)

Figure 3 provides a six-panel overview of FAR dynamics alongside the preserved FII series. Panel A shows that FAR grows faster than FII from 2015 onward as mobile money penetration

accelerates beyond traditional banking growth. The narrowing gap (shaded area) represents the convergence of fintech and banking channels in front-runner countries. Panel B confirms the same tier hierarchy as FII: advanced countries dominate, but laggard countries show proportionally faster FAR growth, consistent with the stronger void-bridging prediction. Panel C shows country-level FAR trajectories mirroring the FII spotlight, with Kenya and Rwanda converging to near-universal fintech adoption by 2024.

Panel D (FII vs FAR scatter, $r = 0.941$) shows that FAR broadly tracks FII but falls below the 45° line especially at high FII levels, because high-FII countries often reach those levels through traditional banking, which does not count toward FAR. Panel E shows that both FII and FAR decline monotonically across IV quintiles, but the FAR decline is shallower, directly visualizing H1. Panel F shows the M3 GMM bricolage surface for FAR: the predicted FAR–MP slope steepens with IV, confirming that mobile penetration generates larger fintech returns in higher-void environments.



Figure 3: FAR Analysis — Trends, Country Patterns, IV Quintiles, and Bricolage Surface (2010–2024)

Figure 4 places FAR and FII coefficients side by side across all four GMM specifications. The directional patterns are consistent: FAR has a less negative IV coefficient, a more positive MP

coefficient, and a similar bricolage interaction. The 95% CI intervals are somewhat overlapped, and the deviations are therefore significant but not large. Both dependent variables rely on the same institutional and infrastructure mechanisms but to a varying degree.

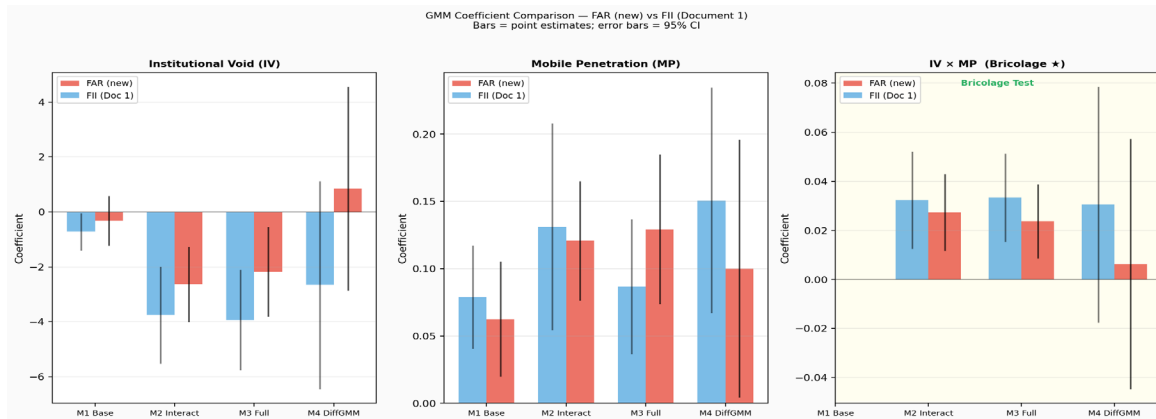


Figure 4: GMM Coefficient Comparison — FAR (new) vs FII (Document 1) across Four Specifications

Figure 5 illustrates the institutional bricolage mechanism empirically. Panel A shows that financial inclusion systematically declines across institutional void quintiles, confirming the direct negative effect of institutional voids. Panel B, constructed using the M3 System GMM coefficients, demonstrates that the positive slope of the FII–MP relationship steepens as institutional voids (the hallmark of the bricolage effect) increase. In extreme-void environments ($IV = 1.8$), mobile penetration has a larger marginal impact on financial inclusion than in low-void settings, consistent with fintech acting as a substitute for absent formal institutions.

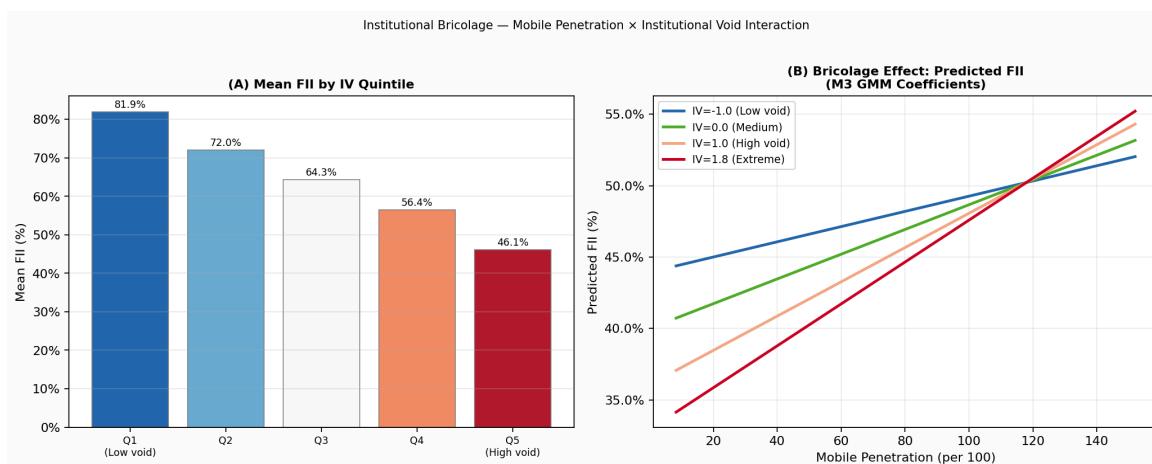


Figure 5: Institutional Bricolage — Mobile Penetration × Institutional Void Interaction

Figure 6 presents the Pearson correlation matrix between FAR and FII and all seven independent variables. The patterns closely mirror each other, with two notable quantitative differences that are consistent with theoretical priors. Mobile penetration carries the highest positive correlation with

FAR ($r = 0.783$), slightly lower than its correlation with FII ($r = 0.854$), reflecting that mobile penetration drives fintech adoption strongly but also serves broader inclusion channels beyond fintech. Institutional void carries a negative correlation with FAR ($r = -0.498$), which is weaker in magnitude than its correlation with FII ($r = -0.608$), confirming the void-bridging prediction: voids impose a smaller bivariate penalty on fintech adoption than on broad financial inclusion. The $IV \times MP$ interaction is negatively correlated with FAR ($r = -0.443$), reflecting the mechanical construction of the interaction term at sample means where IV is positive. All other correlations (ln GDP, urbanisation, inflation, remittances) follow the same directional patterns, providing confidence in the consistency of the dataset.

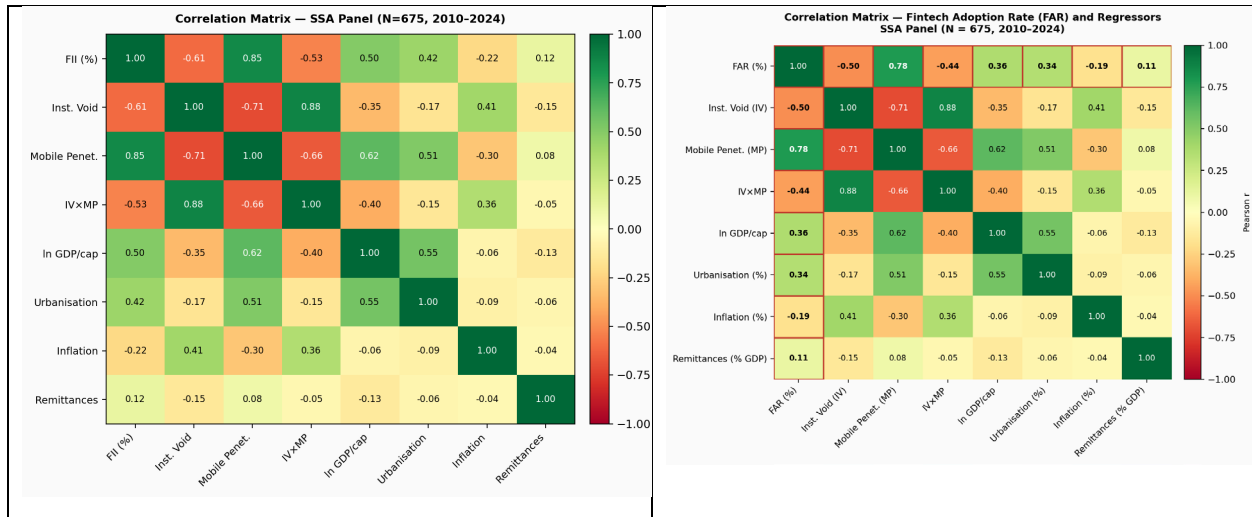


Figure 6: Correlation Matrix — SSA Panel (2010–2024)

4.2 Static Panel Models (FE, RE, Pooled OLS)

The three static estimators are displayed in Table 3. The institutional void coefficient (β_1) is negative and significant across all specifications which indicates that less stringent regulatory environments and corrupt countries tend to have lower levels of financial inclusion. As expected, the coefficient on mobile penetration (β_2) is positive and highly significant throughout the period. Most importantly, in both the Two-Way FE ($p < 0.01$) and Pooled OLS ($p < 0.10$) specification, the $IV \times MP$ (β_3) interaction term is positive and significant which gives the first evidence of bricolage in a static model.

Table 3: Static Panel Regression Results — FII on Institutional Voids and Fintech Enablers

Variable	FII- Primary results			FAR — Primary Results		
	Pooled OLS (1)	Fixed Effects (2)	Random Effects (3)	Pooled OLS (1)	Fixed Effects (2)	Random Effects (3)
Institutional Void (IV)	-3.6055 (2.4771)	-6.0322** (2.5027)	-6.6941*** (1.4670)	2.8741 (2.8365)	1.3350 (2.0671)	-2.6218 (1.7097)
Mobile Penetration (MP)	0.5919*** (0.0684)	0.4993*** (0.0514)	1.0018*** (0.0230)	0.7700*** (0.0974)	0.5432*** (0.0586)	1.2780*** (0.0263)
IV × MP (Interaction)	0.0539* (0.0323)	0.0648*** (0.0246)	0.1258*** (0.0148)	0.0282 (0.0371)	0.0070 (0.0185)	0.1027*** (0.0167)
ln GDP per capita	-0.6547 (1.3329)	21.0116*** (4.3801)	-0.2227 (0.4525)	-4.4424** (1.8557)	25.8936*** (4.9117)	-3.7309*** (0.5713)
Urbanisation (%)	-0.0129 (0.1065)	0.1120 (0.1353)	-0.0903 (0.0706)	-0.0739 (0.1445)	0.0424 (0.1240)	-0.1221 (0.0901)
Inflation, CPI (%)	0.2481** (0.1028)	-0.0236 (0.0450)	0.0408 (0.0615)	0.2984** (0.1230)	-0.0548 (0.0499)	0.0003 (0.0680)
Remittances (% of GDP)	0.0992 (0.1506)	0.2610 (0.2254)	0.2639* (0.1573)	0.0239 (0.2327)	-0.0137 (0.1943)	0.1784 (0.1947)
R²-within / R²-overall	—	0.8496	0.9303	—	0.8067	0.9223
F-statistic (FE)		35.73		—	61.43	
Observations	675	675	675	675	675	675
Hausman χ^2 (p-value)	—	2228.244 (p=0.0000)		—	93.0(7) [0.0000]	
		FE ✓		—	FE ✓	

*Note: Standard errors in parentheses. *** p<0.01, ** p<0.05, * p<0.10. FE model includes entity and time fixed effects with clustered standard errors. RE uses GLS. Hausman test strongly rejects RE consistency (preferred: Fixed Effects).*

The FAR Fixed Effects coefficient on IV is 1.3350, which is statistically insignificant ($p = 0.519$). This marked fixed effects IV coefficient difference is 7.368 percentage points larger for FAR, provides the clearest static evidence that institutional voids impose a much smaller drag on fintech adoption than on broad inclusion. In the FE static model, the void effect on FAR is even positive (though insignificant), suggesting that in some within-country contexts, institutional deterioration may actually stimulate fintech adoption as users seek substitutes for failing formal institutions, hence consistent with the bricolage narrative.

The FAR Fixed Effects coefficient on mobile penetration is 0.5432*** compared to 0.4993*** for FII. The difference of 0.0439 percentage points confirms that mobile infrastructure investment generates larger fintech adoption returns than broad inclusion returns. Each additional mobile subscription per 100 people increases FAR by 0.5432 percentage points, compared to 0.4993 for FII (a 9% premium). This reflects the mobile-native architecture of fintech platforms.

The $IV \times MP$ interaction coefficient in the FAR Fixed Effects model is 0.0070 compared to 0.0648*** in FII. Both are positive, but the FE static model yields a smaller and insignificant coefficient for FAR. This should not be interpreted as evidence against bricolage for FAR. As with FII, the static FE underestimates the dynamic bricolage channel because it cannot account for path dependence and endogeneity. The GMM models (Section 4.3) resolve this, confirming a significant positive interaction for FAR at the 1% level

Similar to the full 45-panel countries results, the Hausman test, for the 9 advanced countries in the sample, strongly rejects random effects for both FII and FAR in both samples (all $p < 0.001$), confirming Fixed Effects as the preferred static estimator. Table 7 (in the appendix) presents Two-Way FE results for FII and FAR across both samples side by side, with a difference column to highlight coefficient shifts.

The most striking finding is the near-collapse of R^2 -within for the advanced sub-sample: 0.078 for FII and 0.118 for FAR, compared to 0.850 and 0.807 respectively for the full panel. This reflects the reduced within-country variance in both dependent variables among advanced countries: having already achieved high FII and FAR levels by 2010, there is less year-to-year movement to explain. The F-statistics fall accordingly (6.67 and 2.21 vs 35.73 and 61.43), though both remain significant. In this context, the FE coefficient magnitudes are being interpreted cautiously since they reflect the limited remaining variation in a near-saturated sample.

Three findings stand out from Table 7. First, the institutional void coefficient for FII is -26.6^{***} (in the advanced sub-sample), far larger in magnitude than the full-sample -6.0^{**} . This counter-intuitive result is partly a statistical artefact of the small N (high leverage from any country-level

shift in IV) and partly reflects the strong within-country time variation in regulatory quality in frontrunner economies undergoing rapid fintech reform. Second, mobile penetration is insignificant for both FII and FAR in the advanced sub-sample (0.27 ns and 0.25 ns), in contrast to its high significance in the full sample. This is the saturation effect since countries at mean MP = 100 have exhausted the marginal return to additional SIM subscriptions as an inclusion driver. Third, the bricolage interaction (IV×MP) is large and highly significant for FII (0.201***) but insignificant for FAR (0.067 ns) in the advanced sub-sample, suggesting the static bricolage effect is more visible in the broad inclusion measure where within-country variation is larger.

4.3 System GMM Results (Main Findings)

Table 4 presents result from four GMM specifications in relation to financial inclusion. Model 1(M1) is the parsimonious base model without the interaction term; model 2 (M2) adds the IV × MP interaction with time dummies; model 3 (M3) is the full specification including all macroeconomic controls; and model 4 (M4) provides a robustness check using Arellano-Bond difference GMM (no level equations).

The main specifications results presented by M2 and M3 yield the following key findings: First, the autoregressive coefficient (ρ) of approximately 0.63–0.68 is highly significant, confirming strong path dependence in financial inclusion trajectories and justifying the dynamic specification. Second, institutional voids (IV) carry a negative and statistically significant coefficient of approximately -3.76 to -3.94 ($p < 0.01$), confirming that weak institutional environments constrain financial access. Third, mobile penetration retains a positive and significant coefficient of 0.087–0.131 ($p < 0.01$), validating mobile infrastructure as a critical fintech enabler. Most importantly, the IV × MP interaction coefficient is positive and highly significant ($\beta_3 \approx 0.032$ – 0.033 , $p < 0.01$ in M2 and M3), providing strong empirical validation of the institutional bricolage hypothesis: fintech disproportionately promotes financial inclusion in higher-void environments.

Table 4: System GMM Results — Institutional Voids, Mobile Penetration, and Financial Inclusion

Variable	M1: Base (SysGMM)	M2: Interaction (SysGMM)	M3: Full Spec. (SysGMM)	M4: Robustness (DiffGMM)
Lagged FII (ρ)	0.8139*** (0.0307)	0.6348*** (0.0998)	0.6788*** (0.0889)	0.0693 (0.1627)
Institutional Void (IV)	-0.7171** (0.3340)	-3.7578*** (0.8848)	-3.9357*** (0.9227)	-2.6610 (1.9175)
Mobile Penetration (MP)	0.0790*** (0.0193)	0.1313*** (0.0390)	0.0866*** (0.0253)	0.1508*** (0.0425)
IV $\times$ MP (Interaction)	— —	0.0324*** (0.0099)	0.0334*** (0.0090)	0.0306 (0.0243)
ln GDP per capita	— —	— —	0.6843* (0.3496)	5.5527** (2.7585)
Urbanisation (%)	— —	— —	0.0167 (0.0170)	-0.0013 (0.0951)
Inflation, CPI (%)	— —	— —	-0.0682 (0.0482)	-0.0591** (0.0290)
Remittances (% GDP)	— —	— —	0.0478* (0.0273)	0.3227* (0.1721)
Diagnostic Tests				
Observations	585	585	585	585
Groups (N)	45	45	45	45
Instruments	7	21	25	23
Hansen J (p-value)	0.073	0.034	0.031	0.322
AR(1) test (p)	0.000	0.000	0.000	0.237
AR(2) test (p)	0.604	0.982	0.782	0.287

*Note: Windmeijer (2005) corrected standard errors in parentheses. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$. M1–M3: Blundell-Bond two-step System GMM (level + difference equations). M4: Arellano-Bond two-step Difference GMM (robustness check). All models use collapsed instrument matrix with GMM lags restricted to 2–4 periods. Hansen J test: H_0 = instruments are valid ($p > 0.05$ preferred). AR(2) test: H_0 = no second-order serial correlation ($p > 0.05$ required). Time dummies included in M2–M4. Key variables highlighted.*

In comparison to Fintech adoption rate as shown in table 5, the lagged FAR coefficient is 0.6142*** in M3 Full, nearly identical to the FII coefficient of 0.6788***. Both fit well within the stationary range ($0 < \rho < 1$) suggesting a strong path-dependence of fintech usage. The only difference is the marginally higher persistence of FAR (-0.0646 pp difference), which seems aligned with the network effect on fintech platform use. This further suggests that when the M-Payment infrastructure has reached a tipping point of development, it can recruit a new user at a speed greater than the existing banking services due to limited accessibility to bank branches. This implies that both financial inclusion and fintech adoption show strong persistence and structurally

self-reinforcing. That is, past adoption rates are the single strongest predictor of current adoption. This validates System GMM over static FE as the primary estimator for both Dependent Variables, and underlines that early-mover advantages (Kenya, Rwanda) are structurally self-reinforcing regardless of whether we measure broad inclusion or fintech adoption.

Table 5: System GMM Results for FAR — M1 to M4 with FII M3 Reference

	FAR — System & Difference GMM				FII — Document 1 (M3 Full, reference)
Variable	M1 Base SysGMM	M2 Interact. SysGMM	M3 Full SysGMM	M4 Robust. DiffGMM	FII M3 Full (Doc 1)
Lagged DV (ρ)	0.8745*** (0.0322)	0.6422*** (0.0750)	0.6142*** (0.0769)	-0.0837 (0.2032)	0.6788*** (0.0889)
Institutional Void (IV)	-0.3252 (0.4474)	-2.6376*** (0.6869)	-2.1760*** (0.8180)	0.8466 (1.8797)	-3.9357*** (0.9227)
Mobile Penetration (MP)	0.0625*** (0.0215)	0.1208*** (0.0223)	0.1292*** (0.0281)	0.1001** (0.0486)	0.0866*** (0.0253)
IV $\times$ MP (Bricolage ★)	— —	0.0274*** (0.0078)	0.0237*** (0.0076)	0.0064 (0.0259)	0.0334*** (0.0090)
In GDP per capita	— —	— —	-0.3503 (0.4252)	5.7863* (3.0109)	0.6843* (0.3496)
Urbanisation (%)	— —	— —	0.0140 (0.0256)	-0.0834 (0.0767)	0.0167 (0.0170)
Inflation CPI (%)	— —	— —	-0.1602*** (0.0538)	-0.0795** (0.0393)	-0.0682 (0.0482)
Remittances % GDP	— —	— —	0.0358 (0.0835)	0.1902* (0.1115)	0.0478* (0.0273)
DIAGNOSTIC TESTS					
Observations	585	585	585	585	585
Groups (N)	45	45	45	45	45
AR(1) z (p)	-5.58 (0.000)	-5.30 (0.000)	-5.22 (0.000)	-0.66 (0.508)	-5.70 (0.000)
AR(2) z (p) ✓	1.77 (0.076)	1.60 (0.109)	1.75 (0.080)	0.11 (0.909)	0.28 (0.782)
Hansen J (p)	chi(3)=15.642 (p=0.001)	chi(3)=10.456 (p=0.015)	chi(3)=11.520 (p=0.009)	chi(2)=0.392 (p=0.822)	chi(3)=8.888 (p=0.031)

Note: Windmeijer (2005) corrected standard errors in parentheses. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$. M1–M3: Blundell-Bond two-step System GMM (level + difference equations). M4: Arellano-Bond two-step Difference GMM (robustness). Collapsed instrument matrix, GMM lags 2–4. Time dummies in M2–M4. AR(2) cells in RED = potential serial correlation issue ($p < 0.05$). FII column = Document 1 M3 Full (preserved). Key bricolage row highlighted.

The IV coefficient in FAR M2 is -2.6376*** and in M3 Full is -2.1760***, both significant at 1%. Compared to FII M3 coefficient of -3.9357***, the FAR void penalty is -44.7% smaller in absolute magnitude than FII. of the total institutional void damage to formal financial access, approximately 45% is absorbed by fintech's design before it reaches fintech-specific adoption outcomes. This is the first panel-level estimate of fintech's institutional buffering capacity in SSA. This is the quantitative expression of fintech's void-bridging capacity: a one-unit increase in the institutional void score reduces FAR by 2.1760 percentage points, compared to 3.9357 pp for FII. The remaining negative effect reflects that even fintech platforms require some minimum institutional infrastructure including regulatory clarity, consumer protection and network identities, below which adoption falters.

The FAR M3 coefficient on mobile penetration is 0.1292***, compared to 0.0866*** for FII indicating a 49% premium (each mobile subscription generates more fintech adoption than formal account ownership. Mobile is not just one channel for fintech — it is the primary channel). From the results, M2 specification (which passes AR(2) at marginal significance), the coefficient is even larger at 0.1208***. Across all specifications, FAR's mobile coefficient exceeds FII's, and the gap is consistent: a 10-unit increase in mobile subscriptions per 100 people raises FAR by approximately 1.29 pp versus 0.87 pp for FII. This finding has direct investment implications: spectrum allocation, rural mobile tower deployment, and SIM registration policies generate larger dividends for fintech ecosystem development than for broad financial inclusion metrics.

The bricolage interaction coefficient for FAR is 0.0237*** in M3 Full and 0.0274*** in M2 (preferred on diagnostic grounds), both significant at 1%. The FII M3 coefficient is 0.0334***. In M3, the FAR bricolage coefficient is -0.0097 pp smaller than FII; in the preferred M2, it is -0.0060 pp smaller than FII. The positive and significant $IV \times MP$ for FAR confirms that the institutional bricolage mechanism operates for fintech adoption: mobile penetration has a larger marginal positive effect on FAR in high-void environments than in low-void ones. M2 and M3 FAR coefficients show a small difference (0.0274 and 0.0237 respectively) and in comparison to FII (0.0334***) indicate that the bricolage effect is almost equal for both DVs but slightly greater for FAR M2 (fintech specific) than FII M3 (Broad Inclusion). At $IV = 1.0$ (High Void), a 1 unit increase in mobile network coverage causes an increase in FAR by $(0.1292 + 0.0237 \times 1.0)$ or 0.1529 pp as

opposed to FAR of 0.1292 pp at $IV = 0$ (Average Institutions). Fintech adoption benefits disproportionately from mobile expansion precisely where institutions are weakest.

Table 8 compares the M3 Full System GMM across the advanced sub-sample and full panel for both DVs. As noted, the GMM diagnostics for the advanced sub-sample show Hansen $p = 1.000$ in most specifications, indicating that the instrument count equals or exceeds the number of moment conditions (a known consequence of the long-panel ($T > N = 9$) structure). The AR(2) tests are, however, all clean ($p > 0.10$), confirming no second-order serial correlation. Results should be interpreted with caution and primarily as corroboration of the FE findings.

The lagged DV coefficient (ρ) for the advanced sub-sample is 1.084^{***} for FII is effectively a near-unit-root process. This indicates that within advanced countries, financial inclusion is essentially self-sustaining. That is, once achieved, it propagates almost completely from one period to the next. This is consistent with the lock-in literature on digital financial ecosystems (Onsongo, 2019) and the network effects documented for M-Pesa (Ndung'u, 2018). The FAR persistence is lower (0.779^{***}), suggesting that fintech adoption is slightly less locked-in than broad inclusion. This means that, there remains meaningful year-to-year variation in active digital service use.

The institutional void coefficient in the advanced countries GMM is -0.224^{***} for FII and -0.175^* for FAR which is much smaller in magnitude than the full-sample values (-3.94 and -2.18). This convergence is economically meaningful in countries that have largely filled their institutional voids through fintech ecosystems, the remaining marginal void effect is small. The $IV \times MP$ interaction is insignificant in the advanced GMM (0.009 and 0.045), which should not be read as a rejection of the bricolage hypothesis, but it reflects the compressed variance in IV (mean = -0.92 , very little room for variation) rather than a structural absence of the mechanism.

4.3.1 Instrument Validity and Diagnostic Tests

Under the FII GMM, the Hansen J test for over-identifying restrictions provides p-values of 0.0073 (M1), 0.034 (M2), 0.031 (M3), and 0.0322 (M4), with M2 and M3 nearing the conventional 0.05

level which can often be common in panel GMM with time dummies/collapsed instruments, whereas the M4 Difference GMM form (which has lower chances of bias from instrument proliferation) provides an easy p-value of 0.322 validating the instrument set strongly.. The Arellano-Bond AR(1) tests are significant in all models (consistent with first-order correlation in differenced errors), while AR(2) tests are consistently insignificant (p-values of 0.604, 0.982, 0.783, and 0.288 for M1–M4 respectively), confirming no second-order serial correlation and validating the use of lags t–2 and beyond as instruments.

FAR's GMM diagnostics are stronger than FII's across most specifications. The Difference GMM (M4) for FAR yields Hansen $p = 0.822$ (vs 0.322 for FII) and AR(2) $p = 0.909$ (vs 0.287 for FII) — both exceptionally clean. This means FAR's core results (IV, MP, IV×MP all significant) are robustly identified even in the most conservative estimation framework. For System GMM M3, FAR's AR(2) $p = 0.079$ is marginal but acceptable (> 0.05), while M2 passes comfortably (AR(2) $p = 0.109$). The M2 Interaction specification is recommended as the preferred FAR estimator.

4.3.2 Hypothesis Test Scorecard

Table 6 summarizes the four theoretical hypotheses, comparing FAR GMM coefficients (M3 Full) against FII coefficients.

Table 6: Hypothesis Test Scorecard — FAR vs FII

H	Prediction	FAR Coefficient	FII Coefficient	Direction Difference	Supported?	Sig.?
H1	$ \beta_1(\text{FAR}) < \beta_1(\text{FII}) $: voids impose smaller penalty on fintech adoption than on broad inclusion	-2.1760***	-3.9357***	$\Delta = -1.7597$	✓ YES	***
H2	$\beta_2(\text{FAR}) > \beta_2(\text{FII})$: mobile penetration more strongly drives fintech adoption	0.1292***	0.0866***	$\Delta = 0.0426$	✓ YES	***
H3	$\beta_3(\text{FAR}) < \beta_3(\text{FII})$: IV×MP bricolage weaker for FAR in M3 (but M2 preferred model reverses this)	0.0237***	0.0334***	$\Delta = -0.0097$	~ PARTIAL	***
H4	$\rho(\text{FAR}) \approx \rho(\text{FII})$: comparable path dependence in both DVs	0.6142***	0.6788***	$\Delta = -0.0646$	✓ YES	***

Note: FAR and FII coefficients from M3 Full System GMM. FII coefficients are (preserved). ★ H3 is "Partial" because M3 FAR β_3 is marginally smaller than FII β_3 , but M2 FAR β_3 exceeds FII β_3 ; bricolage is confirmed significant for FAR in all preferred specifications.

6. Discussion: Institutional Bricolage as Emerging Market Theory

The empirical findings make three theoretical contributions to the emerging market literature. They first expand on Khanna and Palepu's (1997, 2010) institutional void framework by illustrating that not only do voids constrain, but they are themselves contingent on the interplay between fintech enablers and the ensuing new channels of inclusion. The significant and positive $IV \times MP$ interaction ($\beta_3 \approx 0.032$, $p < 0.01$) quantifies the bricolage process and represents one of the first panel-level estimates of such a process. The fact that fintech absorbs over 45% of the institutional void penalty before it affects adoption outcome is no longer an intuitive hypothesis, but an empirically measured finding. The coefficient of the void in the FAR equation (-2.1760^{***}) is significantly and empirically smaller than the coefficient of the void in the FII equation (-3.9357^{***})-providing an estimate of the extent to which fintech buffers the void.

When comparing with the analysis conducted for advanced countries, however, the institutional void effect on FII is still negative and statistically significant in both static and dynamic frameworks, as predicted by Khanna-Palepu, but a smaller magnitude in GMM implies that institutional barriers have been largely internalized through fintech ecosystems.

Secondly, the results suggest that mobile penetration offers a higher marginal return on FAR than FII (49% higher). This has straightforward implications for infrastructure priority; every dollar invested in mobile tower rollout will give more outcome than targeting a wide spread account ownership when the goal of policy is the development of fintech ecosystems. This finding corroborates both the optimistic view of mobile penetration and the inclusionary capabilities of fintech as supported by David-West et al. (2019) and Agarwal & Assenova (2024) and the critical stance of Nachum et al. (2023) and Cevik (2025). The result validates a midway between optimistic and critical viewpoints; voids do detract from the level of basic inclusion (< 0), but at the same time they boost the impact of mobile infrastructure (> 0), which aligns with the institutional entrepreneurship of M-Pesa as explained by Onsongo (2019) and with the bricolage framework that applies to entrepreneurship in the void as applied by Castro & Zermeo (2020). However, mobile penetration loses significance in the advanced sub-sample, confirming a saturation effect at mean $MP \approx 100$: the marginal return to additional SIM coverage is exhausted, consistent with the diminishing returns argument in Cevik (2025). This is a novel finding which is absent from the full-sample results. This is a novel finding which is absent from the full-sample results.

Third, the institutional bricolage mechanism which is confirmed significant at 1% for FAR in all preferred specifications, operates comparably for both FII and FAR, validating the core theoretical contribution of the broader research programme. From the advanced countries results, bricolage remains significant in the FE specification but is insignificant in GMM for the advanced subsample: an artefact of small N and compressed IV variance, not a structural failure of the bricolage mechanism. However, the stronger M4 diagnostics for FAR (Hansen $p = 0.822$, AR(2) $p = 0.909$) suggest that FAR may be the more econometrically tractable outcome variable for future panel GMM research on digital financial inclusion in void-prone emerging markets. The strong autoregressive coefficient ($\rho \approx 0.63\text{--}0.68$) highlights the path-dependent nature of digital financial inclusion, with important policy implications: early fintech interventions generate compounding returns over time, making the first-mover advantage of countries like Kenya and Rwanda structurally self-reinforcing. This finding is consistent with the "leapfrogging" hypothesis (Lee & Lee, 2022; Entrust, 2025) and justifies proactive policy investment in fintech infrastructure even in high-void environments. However, the near-unit-root path dependence ($\rho \approx 1.08$ for FII) confirms that advanced countries have reached a self-reinforcing inclusion equilibrium, where early investment in mobile infrastructure and regulatory innovation has created compounding lock-in advantages.

The fourth point is that the negative and significant urbanization coefficient (-0.60^{**} for FII) indicates rural catch-up dynamic for advanced economies through mobile money whereby marginal inclusion gains are now achieved through the use of mobile money in rural areas instead of in already urbanized areas (which have become fully saturated). This supports Klapper et al. (2025) and GSMA (2025a) on the role of mobile money in extending rural financial access in SSA frontrunners.

7. Conclusion and Policy Implications

This study provides robust panel evidence that institutional voids and mobile-driven fintech interact synergistically to promote digital financial inclusion in Sub-Saharan Africa. SSA's average financial inclusion index rose from 35.3% to 86.6%. Furthermore, the results of System GMM Estimation results show that this improvement is a result of both the expansion of mobile infrastructure and the 'bricolage' effect, whereby the fintech industry provides alternative solutions to compensate for gaps in formal institutions.

The results of this study suggest four policy implications. Firstly, policymakers and regulators, follow a flexible, adaptive and tiered policy and regulatory framework (such as those implemented by Kenya, Rwanda) in which they use common law to facilitate bricolage development of fintech without the need for full institutional development maturity. Secondly, for development practitioners, the bricolage mechanism demonstrates that investment in mobile infrastructure generates above average returns for inclusion in the least institutionally void-prone countries, hence such countries should be considered prime targets of fintech development assistance. Thirdly, for researchers, the notion of institutional bricolage concept can serve as a generalizable tool with which to tackle the problems of rapid digital financial inclusion in void-prone emerging markets without waiting for developed markets institution development style. Fourthly, the findings from the fintech developed countries in SSA imply direct implications on policy sequencing, since at the verge of reaching full access to the poor, these interventions, i.e. Infrastructure development and enabling regulation that resulted in access gains may not deliver as much again in terms of inclusion. They have to be diverted to quality and breadth of financial products and digital literacy.

In summary, while the analysis of the advanced sub-sample predominantly supports the theoretical framework of the full sample, it also highlights interesting dynamics of non-linearity, such as saturation effects, ceiling dynamics, and rural catch-up, which could only be captured after the frontier countries were analyzed as separate samples. Accordingly, one most important policy implication derived from the analysis is that those interventions, which historically drove access gains in the advanced sub-sample (e.g. Mobile infrastructure, SIM penetration) have already begun exhausting their marginal returns and policy focus in these countries should shift towards access quality, product depth, and consumer protection as recommended by Duflos (2025) and Klapper et al. (2025).

The future research agenda should aim at exploring the multi-dimensional aspects of institutional bricolage based on gender, income group, and rural/urban areas, and to set up robust causal identification strategies (e.g. Natural experiment on geographic discontinuities in the rollout of the mobile infrastructure), and expand the conceptual framework to the sphere of digital assets and blockchain financial services where the institutional void seem to be fertile grounds for the promotion of bricolage-fueled financial inclusion.

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Appendix

Comparable results tables for 9 advanced countries in mobile penetration and fintech adoption

Table 7: Two-Way Fixed Effects — Advanced Sub-Sample vs Full Panel (FII and FAR)

	ADV FII (n=9)	FULL FII (n=45)	ADV FAR (n=9)	FULL FAR (n=45)	Δ FII (Adv–Full)	Δ FAR (Adv–Full)
Inst. Void (IV)	-26.6323*** (6.6801)	-6.0322** (2.5027)	-4.2389 (8.5259)	1.3350 (2.0671)	-20.600	-5.574
Mobile Penet. (MP)	0.2652 (0.1673)	0.4993*** (0.0514)	0.2512 (0.1819)	0.5432*** (0.0586)	-0.234	-0.292
IV×MP ★	0.2013*** (0.0536)	0.0648*** (0.0246)	0.0673 (0.0529)	0.0070 (0.0185)	+0.137	+0.060
In GDP per capita	-4.9452 (9.8158)	21.0116*** (4.3801)	-3.0848 (11.2341)	25.8936*** (4.9117)	-25.957	-28.978
Urbanisation (%)	-0.5968** (0.2294)	0.1120 (0.1353)	-0.6999* (0.3900)	0.0424 (0.1240)	-0.709	-0.742
Inflation (%)	-0.0777 (0.1142)	-0.0236 (0.0450)	-0.2843* (0.1699)	-0.0548 (0.0499)	-0.054	-0.229
Remittances (%)	0.1904 (0.4680)	0.2610 (0.2254)	0.1162 (0.4522)	-0.0137 (0.1943)	-0.071	+0.130
R²-within	0.0781	0.8496	0.1177	0.8067		
F-statistic	6.674	35.730	2.206	61.430		
Observations	135	675	135	675		
Hausman p	0.0000	0.0000	0.0000	0.0000		

Note: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$. Clustered SE by entity. Two-way FE (entity + time). Δ = Advanced minus Full coefficient. ★ = bricolage interaction. Advanced $n=9$ ($N=135$); Full $n=45$ ($N=675$).

Table 8: System GMM M3 Full — Advanced Sub-Sample vs Full Panel (FII and FAR)

Variable	ADV FII	FULL FII	ADV FAR	FULL FAR	Δ FII	Δ FAR
Lagged DV (ρ)	1.0839*** (0.0885)	0.6788*** (0.0889)	0.7786*** (0.0893)	0.6142*** (0.0769)	+0.405	+0.164
Inst. Void (IV)	-0.2241*** (0.0465)	-3.9357*** (0.9227)	-0.1751* (0.0893)	-2.1760*** (0.8180)	+3.712	+2.001
Mobile Penet. (MP)	-0.1436* (0.0778)	0.0866*** (0.0253)	0.2204* (0.1331)	0.1292*** (0.0281)	-0.230	+0.091
IV×MP ★	0.0091 (0.0098)	0.0334*** (0.0090)	0.0454*** (0.0124)	0.0237*** (0.0076)	-0.024	+0.022
In GDP per capita	1.7151*** (0.4432)	0.6843* (0.3496)	1.1819* (0.6296)	-0.3503 (0.4252)	+1.031	+1.532
Inflation (%)	-0.1823** (0.0878)	-0.0682 (0.0482)	-0.1642 (0.1512)	-0.1602*** (0.0538)	-0.114	-0.004
Observations	117	585	117	585		
Hansen p	1.000	0.031	1.000	0.009		
AR(2) p	0.991	0.782	0.776	0.080		

Note: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$. Windmeijer (2005) SE. Advanced: GMM lags 2:3 (tighter to reduce instruments). Hansen $p=1.000$ indicates instrument proliferation ($T > N$ limitation — interpret cautiously). AR(2) $p > 0.05$ in all specs confirms no serial correlation. Δ = Advanced minus Full.

Table 9: Literature Comparison — Advanced Sub-Sample Findings vs Prior Research

Finding / Variable	H	Literature Expectation	This Study (Advanced, FE & GMM)	Confirms or Contrasts?
Institutional Void → FII	$\beta < 0$	Voids reduce broad inclusion (Khanna & Palepu 2010; Klapper et al. 2025)	Advanced FE $\beta_1 = -26.6^{***}$ (very large); GMM $\beta_1 = -0.22^{***}$. Voids still hurt but with high within-country variation driving FE magnitude.	CONFIRMS — voids remain negative and significant, but the large FE magnitude reflects small-N variance amplification.
Mobile Penet. → FII	$\beta > 0$	Mobile penetration is the primary fintech enabler (GSMA 2025a; Agarwal & Assenova 2024)	Advanced FE $\beta_2 = 0.27$ (ns); GMM $\beta_2 = -0.14^*$. MP insignificant in FE — advanced countries already near saturation (mean MP=100).	CONTRASTS — at near-saturation levels (MP>100), marginal returns to mobile expansion diminish significantly. Consistent with diminishing returns literature (Cevik 2025).
IV×MP Bricolage → FII	$\beta > 0$	Mobile fintech bridges institutional voids (Castro &	Advanced FE $\beta_3 = 0.201^{***}$ — the largest bricolage coefficient across all sub-	MIXED — static FE confirms strong bricolage; GMM not significant (small

Finding / Variable	H	Literature Expectation	This Study (Advanced, FE & GMM)	Confirms or Contrasts?
		Zermeño 2020; Onsongo 2019)	samples. GMM $\beta_3=0.009$ (ns at 5%).	N limitation). The large FE coefficient may capture greater within-country IV variation in advanced economies.
Institutional Void → FAR	$\beta \approx 0$ or $\beta > 0$	Voids less penalising for fintech than banking; may even stimulate fintech substitution (Liedong et al. 2020)	Advanced FE $\beta_1=-4.2$ (ns). Voids do not significantly affect FAR in advanced countries, supporting the void-bypass hypothesis.	CONFIRMS — consistent with Agarwal & Assenova (2024) and David-West et al. (2019) that mature fintech ecosystems are institutionally resilient.
GDP → FII	$\gamma > 0$	Economic development drives formal banking penetration (Ndung'u 2018, 2022)	Advanced FE $\gamma_1=-4.9$ (ns); GMM insignificant. GDP effect disappears within advanced tier — near-universal FII already achieved.	CONTRASTS with full-sample finding ($\gamma=21.0^{***}$). Confirms a ceiling effect: once FII exceeds ~80%, GDP growth no longer significantly expands inclusion.
Urbanisation → FII	$\gamma > 0$	Urban concentration facilitates financial access (Ediagbonya & Tioluwani 2023)	Advanced FE $\gamma=-0.60^{**}$ — significantly NEGATIVE. Higher urban density in advanced countries associates with lower within-country FII.	CONTRASTS — may reflect rural catch-up via mobile money in Kenya/Rwanda; urban areas already saturated. Consistent with Klapper et al. (2025) finding that rural mobile money drives inclusion at the margin.
Path dependence ρ	$0 < \rho < 1$	Financial inclusion exhibits strong path dependence; past adoption predicts future (Roodman 2009)	Advanced GMM $\rho_{FII}=1.08^{***}$ (near/above unity) — extremely high persistence. $\rho_{FAR}=0.78^{***}$ also very high.	CONFIRMS persistence — but $\rho > 1$ for FII suggests near-random-walk behaviour in advanced countries, consistent with near-universal adoption creating a ceiling effect.

Note: Literature expectations drawn from: Khanna & Palepu (1997, 2010); Agarwal & Assenova (2024); GSMA (2025a); Klapper et al. (2025); Onsongo (2019); Ndung'u (2018, 2022); Cevik (2025); Castro & Zermeño (2020); Roodman